

MATH20802: STATISTICAL METHODS
SECOND SEMESTER
ANSWERS TO THE IN CLASS TEST

ANSWERS TO QUESTION 1

Let X be a normal random variable with mean μ and standard deviation σ . Let $Y = \exp(2X)$; that is, Y is a log-normal random variable with parameters 2μ and 2σ .

(i) If $X \sim N(\mu, \sigma^2)$ then

$$\begin{aligned}
 M_X(t) &= E[\exp(tX)] \\
 &= \int_{-\infty}^{\infty} \exp(tx) \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} dx \\
 &= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2} + tx\right\} dx \\
 &= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} \exp\left\{-\frac{x^2 - 2\mu x - 2\sigma^2 tx + \mu^2}{2\sigma^2}\right\} dx \\
 &= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} \exp\left\{-\frac{(x - \mu - \sigma^2 t)^2 + \mu^2 - (\mu + \sigma^2 t)^2}{2\sigma^2}\right\} dx \\
 &= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} \exp\left\{-\frac{(x - \mu - \sigma^2 t)^2 - 2\mu\sigma^2 t - \sigma^4 t^2}{2\sigma^2}\right\} dx \\
 &= \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right) \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} \exp\left\{-\frac{(x - \mu - \sigma t)^2}{2\sigma^2}\right\} dx \\
 &= \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right).
 \end{aligned}$$

(ii) We have $E(Y) = M_X(2)$ since $E(Y) = E(\exp(2X))$. We have $E(Y^2) = M_X(4)$ since $E(Y^2) = E(\exp(4X))$.

(iii) From (ii), we have

$$E(Y) = M_X(2) = \exp\left(2\mu + 2\sigma^2\right)$$

and

$$E(Y^2) = M_X(4) = \exp\left(4\mu + 8\sigma^2\right)$$

So,

$$\text{Var}(Y) = \exp\left(4\mu + 8\sigma^2\right) - \exp\left(4\mu + 4\sigma^2\right).$$

(iv) We have

$$M_{X_1+X_2}(t) = E[\exp(tX_1 + tX_2)]$$

$$\begin{aligned} &= E[\exp(tX_1)] E[\exp(tX_2)] \\ &= \exp\left(\mu_1 t + \frac{\sigma_1^2 t^2}{2}\right) \exp\left(\mu_2 t + \frac{\sigma_2^2 t^2}{2}\right) \\ &= \exp\left((\mu_1 + \mu_2) t + \frac{(\sigma_1^2 + \sigma_2^2) t^2}{2}\right). \end{aligned}$$

(v) It follows from (iv) that $2X_1 + 2X_2$ is a normal random variable with mean $2\mu_1 + 2\mu_2$ and variance $4\sigma_1^2 + 4\sigma_2^2$. Hence, $\exp(2X_1 + 2X_2)$ is a log-normal random variable with parameter $2\mu_1 + 2\mu_2$ and $2\sqrt{\sigma_1^2 + \sigma_2^2}$.

ANSWERS TO QUESTION 2

Suppose X_1, X_2, \dots, X_n is a random sample from $N(c\mu, c\sigma^2)$, where both μ and σ^2 are unknown parameters and c is a fixed known constant.

(i) The joint likelihood function of μ and σ^2 is

$$\begin{aligned} L(\mu, \sigma^2) &= \prod_{i=1}^n \left\{ \frac{1}{\sqrt{2\pi}\sqrt{c\sigma}} \exp \left[-\frac{(X_i - c\mu)^2}{2c\sigma^2} \right] \right\} \\ &= \frac{1}{(2\pi)^{n/2} c^{n/2} \sigma^n} \exp \left[-\frac{1}{2c\sigma^2} \sum_{i=1}^n (X_i - c\mu)^2 \right]. \end{aligned}$$

The joint log likelihood function of μ and σ^2 is

$$\log L(\mu, \sigma^2) = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log c - n \log \sigma - \frac{1}{2c\sigma^2} \sum_{i=1}^n (X_i - c\mu)^2.$$

The first order partial derivatives of this with respect to μ and σ are

$$\frac{\partial \log L}{\partial \mu} = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - c\mu) = \frac{1}{\sigma^2} \left[\left(\sum_{i=1}^n X_i \right) - nc\mu \right] \quad (1)$$

and

$$\frac{\partial \log L}{\partial \sigma} = -\frac{n}{\sigma} + \frac{1}{c\sigma^3} \sum_{i=1}^n (X_i - c\mu)^2, \quad (2)$$

respectively.

(ii) Using equation (1), one can see that the solution of $\partial \log L / \partial \mu = 0$ is $\mu = \bar{X}/c = (1/(nc)) \sum_{i=1}^n X_i$.

(iii) Using equation (2), one can see that the solution of $\partial \log L / \partial \sigma = 0$ is $\sigma^2 = (1/(nc)) \sum_{i=1}^n (X_i - \bar{X})^2$.

(iv) The mle, $\hat{\mu}$, is an unbiased and consistent estimator for μ since

$$\begin{aligned} E(\hat{\mu}) &= E\left(\frac{1}{nc} \sum_{i=1}^n X_i\right) \\ &= \frac{1}{nc} \sum_{i=1}^n E(X_i) \\ &= \frac{1}{nc} \sum_{i=1}^n c\mu \\ &= \mu \end{aligned}$$

and

$$\begin{aligned} Var(\hat{\mu}) &= Var\left(\frac{1}{nc} \sum_{i=1}^n X_i\right) \\ &= \frac{1}{n^2 c^2} \sum_{i=1}^n Var(X_i) \end{aligned}$$

$$\begin{aligned}
&= \frac{1}{n^2 c^2} \sum_{i=1}^n c \sigma^2 \\
&= \frac{\sigma^2}{nc}.
\end{aligned}$$

(v) The mle, $\widehat{\sigma^2}$, is a biased and consistent estimator for σ^2 since

$$\begin{aligned}
E(\widehat{\sigma^2}) &= E\left[\frac{1}{nc} \sum_{i=1}^n (X_i - \bar{X})^2\right] \\
&= E\left[\frac{n-1}{nc} S^2\right] \\
&= \frac{\sigma^2}{n} E\left[\frac{n-1}{c\sigma^2} S^2\right] \\
&= \frac{\sigma^2}{n} E[\chi_{n-1}^2] \\
&= \frac{(n-1)\sigma^2}{n}
\end{aligned}$$

and

$$\begin{aligned}
Var(\widehat{\sigma^2}) &= Var\left[\frac{1}{nc} \sum_{i=1}^n (X_i - \bar{X})^2\right] \\
&= Var\left[\frac{n-1}{nc} S^2\right] \\
&= \frac{\sigma^4}{n^2} Var\left[\frac{n-1}{c\sigma^2} S^2\right] \\
&= \frac{\sigma^4}{n^2} Var[\chi_{n-1}^2] \\
&= \frac{2(n-1)\sigma^4}{n^2}.
\end{aligned}$$

Note that we have used the fact $(n-1)S^2/(c\sigma^2) \sim \chi_{n-1}^2$. Furthermore, $S^2 = (1/(n-1)) \sum_{i=1}^n (X_i - \bar{X})^2$ denotes the sample variance.