

**MATH38181: Extreme values and financial risk**  
**Semester 1**

**Formulas to remember for the final exam in January 2019**

**Extremal type theorem:** Suppose  $X_1, X_2, \dots$  are independent and identically distributed (iid) random variables with common cumulative distribution function (cdf)  $F$ . Let  $M_n = \max\{X_1, \dots, X_n\}$  denote the maximum of the first  $n$  random variables and let  $w(F) = \sup\{x : F(x) < 1\}$  denote the upper end point of  $F$ . If there are norming constants  $a_n > 0$ ,  $b_n$  and a nondegenerate  $G$  such that the cdf of a normalized version of  $M_n$  converges to  $G$ , i.e.

$$\Pr\left(\frac{M_n - b_n}{a_n} \leq x\right) = F^n(a_n x + b_n) \rightarrow G(x)$$

as  $n \rightarrow \infty$  then  $G$  must be of the same type as (cdfs  $G$  and  $G^*$  are of the same type if  $G^*(x) = G(ax + b)$  for some  $a > 0$ ,  $b$  and all  $x$ ) as one of the following three classes:

$$\begin{aligned} I & : \Lambda(x) = \exp\{-\exp(-x)\}, \quad x \in \mathfrak{R}; \\ II & : \Phi_\alpha(x) = \begin{cases} 0 & \text{if } x < 0, \\ \exp\{-x^{-\alpha}\} & \text{if } x \geq 0 \end{cases} \\ & \quad \text{for some } \alpha > 0; \\ III & : \Psi_\alpha(x) = \begin{cases} \exp\{-(-x)^\alpha\} & \text{if } x < 0, \\ 1 & \text{if } x \geq 0 \end{cases} \\ & \quad \text{for some } \alpha > 0. \end{aligned}$$

**Necessary and sufficient conditions for the three extreme value distributions:**

$$\begin{aligned} I & : \exists \gamma(t) > 0 \text{ s.t. } \lim_{t \uparrow w(F)} \frac{1 - F(t + x\gamma(t))}{1 - F(t)} = \exp(-x), \quad x \in \mathfrak{R}, \\ II & : w(F) = \infty \text{ and } \lim_{t \uparrow \infty} \frac{1 - F(tx)}{1 - F(t)} = x^{-\alpha}, \quad x > 0, \\ III & : w(F) < \infty \text{ and } \lim_{t \downarrow 0} \frac{1 - F(w(F) - tx)}{1 - F(w(F) - t)} = x^\alpha, \quad x > 0. \end{aligned}$$

**Necessary and sufficient conditions for  $(M_n - b_n)/a_n$  to have a non-degenerate limiting distribution:**

$$\frac{\Pr(X = k)}{1 - F(k - 1)} \rightarrow 0$$

as  $k \rightarrow w(F)$  if  $F$  is discrete.

**Definition of  $\text{VaR}_p(X)$  if  $X$  is an absolutely continuous random variable:**

$$\text{VaR}_p(X) = F^{-1}(p).$$

**Definition of  $ES_p(X)$  if  $X$  is an absolutely continuous random variable:**

$$ES_p(X) = \frac{1}{p} \int_0^p F^{-1}(v) dv.$$

**The definition of a moment generating function:  $M_X(t) = E[\exp(tX)]$ .**

**The probability mass and cumulative distribution functions of a Geometric( $\theta$ ) random variable are:**

$$p(k) = \theta(1 - \theta)^{k-1}$$

and

$$F(k) = 1 - (1 - \theta)^k$$

for  $\theta > 0$  and  $k = 1, 2, \dots$

**Definition of beta function:**

$$B(a, b) = \int_0^1 t^{a-1}(1-t)^{b-1} dt$$

for  $a > 0$  and  $b > 0$ .

**Definition of incomplete beta function:**

$$B_x(a, b) = \int_0^x t^{a-1}(1-t)^{b-1} dt$$

for  $a > 0$ ,  $b > 0$  and  $0 < x < 1$ .

**Definition of incomplete beta function ratio:**

$$I_x(a, b) = B_x(a, b)/B(a, b)$$

for  $a > 0$ ,  $b > 0$  and  $0 < x < 1$ .

**L'Hôpital's rule:**

$$\lim_{x \rightarrow c} \frac{f_1(x)}{f_2(x)} = \lim_{x \rightarrow c} \frac{f_1'(x)}{f_2'(x)}$$

if  $\lim_{x \rightarrow c} f_1(x) = \lim_{x \rightarrow c} f_2(x) = 0$  or  $\pm\infty$ .

**The fact that:  $(1-x)^a \approx 1-ax$  for  $x$  close to zero.**

**Joint survival function of  $X_1, X_2, \dots, X_k$  is defined by**

$$\bar{F}(x_1, x_2, \dots, x_k) = \Pr(X_1 > x_1, X_2 > x_2, \dots, X_k > x_k).$$