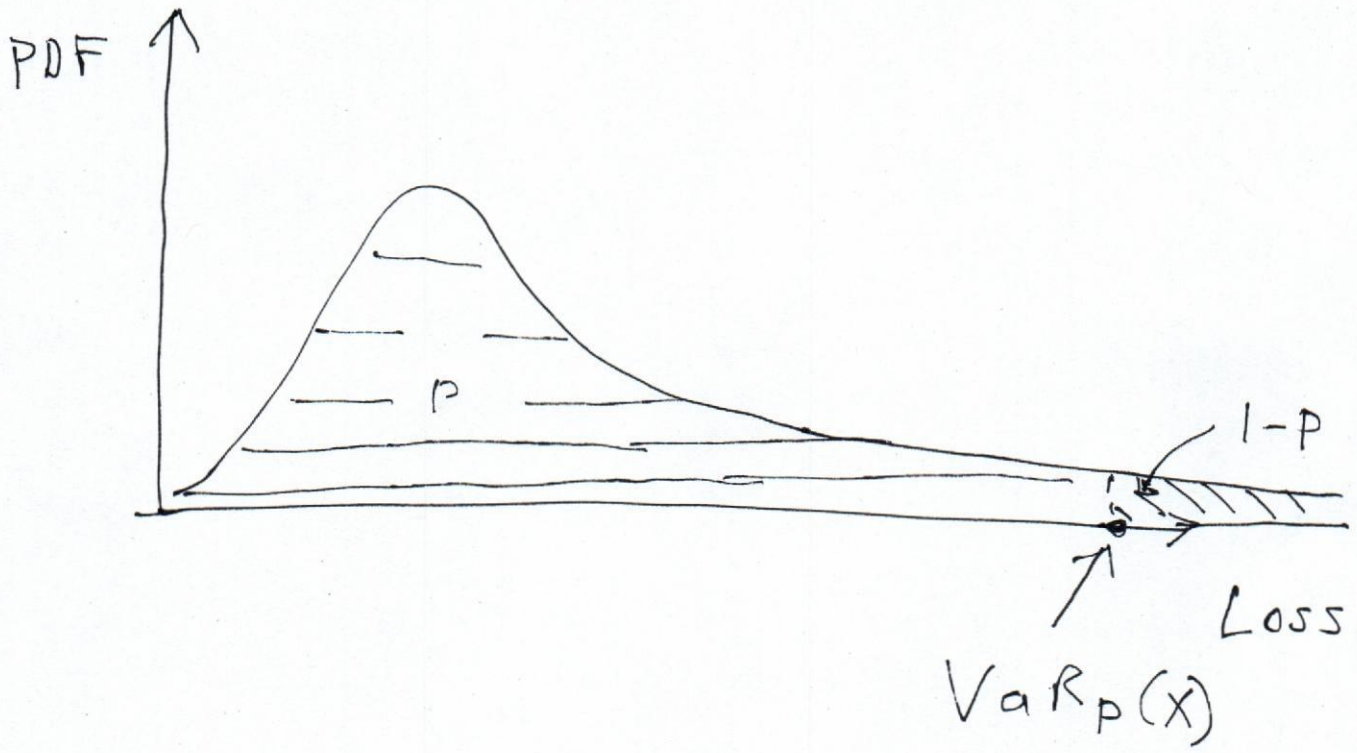


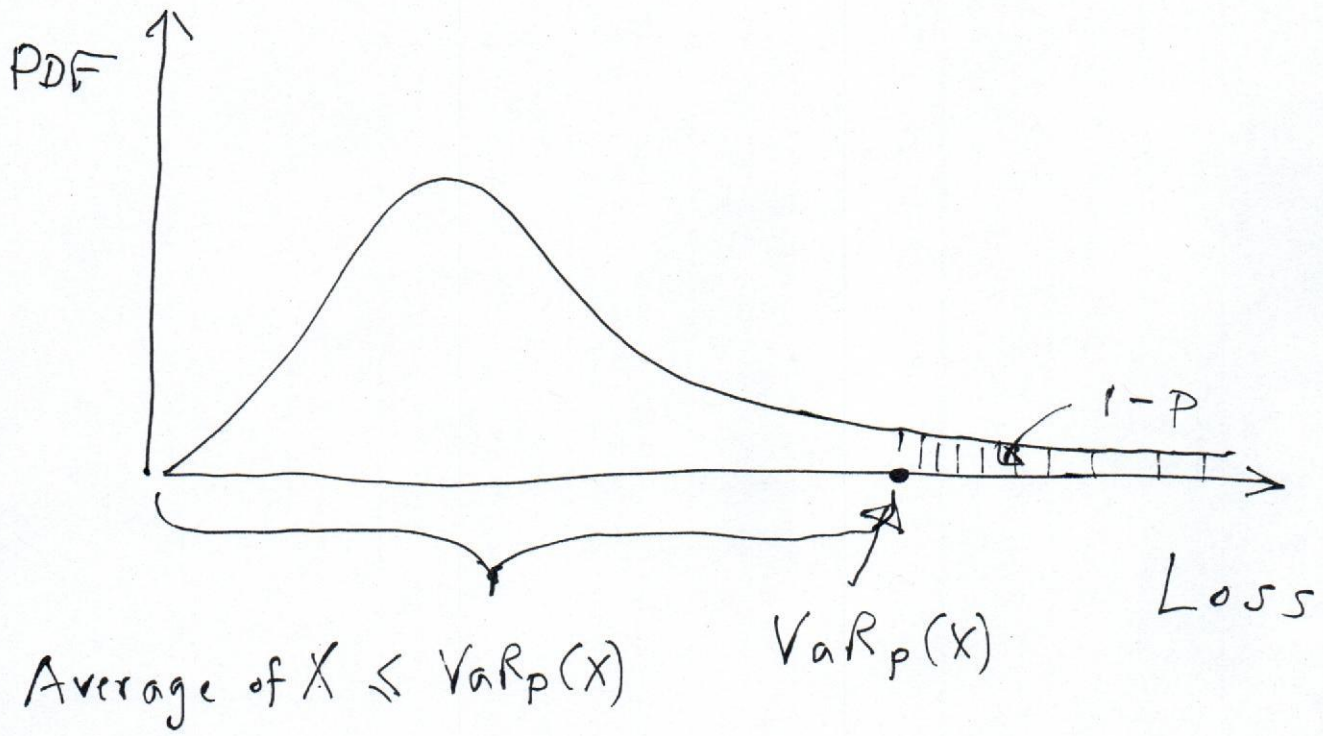
**LECTURE**

**6 NOVEMBER**

**9:00-10:00AM**

**MATH3/4/68181**





$$\parallel$$
$$ES_p(X)$$

Ex 1

Suppose  $X \sim N(\mu, \sigma^2)$

$$F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right) = p$$

$$\Rightarrow \frac{x-\mu}{\sigma} = \Phi^{-1}(p)$$

$$\Rightarrow x = \mu + \sigma \Phi^{-1}(p)$$

$$\Rightarrow \text{Var}_p(X) = \mu + \sigma \Phi^{-1}(p)$$

$$ES_p(X) = \frac{1}{p} \int_0^p [\mu + \sigma \Phi^{-1}(t)] dt$$

$$= \mu + \frac{\sigma}{p} \int_0^p \Phi^{-1}(t) dt$$

Ex 2

$$X \sim \text{Uni} [a, b]$$

$$F(x) = \frac{x-a}{b-a} = p$$

$$\Rightarrow x = a + p(b-a)$$

$$\Rightarrow \text{VaR}_p(X) = a + p(b-a)$$

$$ES_p(X) = \frac{1}{p} \int_0^p [a + t(b-a)] dt$$

$$= \frac{1}{p} \left[ at + \frac{t^2}{2}(b-a) \right]_0^p$$

$$= a + \frac{p}{2}(b-a)$$

Ex 3

Suppose  $X$  has CDF  $F(x) = x^a$ ,  
 $0 < x < 1$

$$F(x) = x^a = p$$

$$\Rightarrow x = p^{\frac{1}{a}}$$

$$\Rightarrow \text{VaR}_p(X) = p^{\frac{1}{a}}$$

$$ES_p(X) = \frac{1}{p} \int_0^p t^{\frac{1}{a}} dt$$

$$= \frac{1}{p} \left[ \frac{t^{\frac{1}{a}+1}}{\frac{1}{a}+1} \right]_0^p$$

$$= \frac{p^{\frac{1}{a}}}{\frac{1}{a}+1} \cdot$$

# Proof of Properties of VaR

Translative property

$$\text{VaR}_p(X+c) = \text{VaR}_p(X) + c \quad (*)$$

Proof. Suppose  $(*)$  holds. Then

$$F_{X+c}^{-1}(p) = F_X^{-1}(p) + c$$

$$\Rightarrow F_{X+c}(F_{X+c}^{-1}(p)) = F_{X+c}(F_X^{-1}(p) + c)$$

$$\Rightarrow p = F_{X+c}(F_X^{-1}(p) + c)$$

$$\Rightarrow p = \Pr(X+c \leq F_X^{-1}(p) + c)$$

$$\Rightarrow p = \Pr(X \leq F_X^{-1}(p))$$

$$\Rightarrow p = F_X(F_X^{-1}(p))$$

$$\Rightarrow p = p.$$

Hence  $(*)$  must hold.

Monotone property

$$X \leq Y \Rightarrow \text{VaR}_p(X) \leq \text{VaR}_p(Y)$$

Proof

$$X \leq Y$$

$$\Rightarrow F_X^{-1}(p) \leq F_Y^{-1}(p) \quad \forall p$$

$$\Rightarrow \text{VaR}_p(X) \leq \text{VaR}_p(Y) \quad \forall p.$$



Positive homogeneity

$$\text{VaR}_p(cX) = c \text{VaR}_p(X) \dots (*)$$

Proof Suppose (\*) holds. Then

$$F_{cX}^{-1}(p) = c F_X^{-1}(p)$$

$$\Rightarrow F_{cX}(F_{cX}^{-1}(p)) = F_{cX}(c F_X^{-1}(p))$$

$$\Rightarrow p = F_{cX}(c F_X^{-1}(p))$$

$$\Rightarrow p = \Pr(\cancel{cX} \leq \cancel{c F_X^{-1}(p)})$$

$$\Rightarrow p = F_X(F_X^{-1}(p))$$

$$\Rightarrow p = p.$$

Hence (\*) must hold.



a) Parametric estimation methods

1) Normal distribution

Suppose  $X \sim N(\mu, \sigma^2)$ . Then

$$\text{Var}_p(X) = \mu + \sigma \Phi^{-1}(p).$$

Suppose  $x_1, x_2, \dots, x_n$  is a random sample on  $X$ . Then the MLEs of  $\mu$  and  $\sigma$  are

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}$$

$$\hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$

Hence, the MLE of  $\text{Var}_p(X)$  is

$$\widehat{\text{Var}}_p(X) = \bar{x} + \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2} \Phi^{-1}(p)$$

Suppose  $\hat{\theta}$  is an estimator of  $\theta$ .

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta$$

$\hat{\theta}$  is unbiased if  $\text{Bias}(\hat{\theta}) = 0$

$$\text{MSE}(\hat{\theta}) = E[(\hat{\theta} - \theta)^2]$$

↑  
Mean Squared Error

$\hat{\theta}$  is consistent if  $\text{MSE}(\hat{\theta}) \rightarrow 0$   
as  $n \rightarrow \infty$ .

$\hat{\theta}$  is asymptotically unbiased  
if  $\text{Bias}(\hat{\theta}) \rightarrow 0$  as  $n \rightarrow \infty$

$\widehat{\text{Var}}_p(X)$  is a biased estimator.

But it is consistent.

## 2) Uniform distribution method

Suppose  $X \sim \text{Uni}[a, b]$ . Then

$$\text{VaR}_p(X) = a + p(b - a)$$

Suppose  $x_1, x_2, \dots, x_n$  is a random sample on  $X$ . Then the MLEs of  $a$  and  $b$  are

$$\hat{a} = \min(x_1, \dots, x_n)$$

$$\hat{b} = \max(x_1, \dots, x_n).$$

Hence the MLE of  $\text{VaR}_p(X)$  is

$$\widehat{\text{VaR}}_p(X) = \min(x_1, \dots, x_n)$$

$$+ p \left[ \max(x_1, \dots, x_n) - \min(x_1, \dots, x_n) \right].$$

Homework :  $\widehat{\text{VaR}}_p(X)$  is a biased estimator but consistent