MATH4/68181: EXTREME VALUES AND FINANCIAL RISK SEMESTER 1 SOLUTIONS TO QUIZ PROBLEM 3

Consider a class of distributions defined by the cdf

$$F(x) = [G(x)]^a \left\{ 1 + \lambda - \lambda \left[G(x) \right]^b \right\}$$

where $a > 0, b > 0, -\infty < \lambda < \infty$ and $G(\cdot)$ is a valid cdf. Assume that F and G have the same upper end points.

First, suppose that G belongs to the max domain of attraction of the Gumbel extreme value distribution. Then, there must exist a strictly positive function say h(t) such that

$$\lim_{t \to w(G)} \frac{1 - G(t + xh(t))}{1 - G(t)} = e^{-x}$$

for every x > 0. But

$$\begin{split} \lim_{t \to w(F)} \frac{1 - F(t + xh(t))}{1 - F(t)} \\ &= \lim_{t \to w(F)} \frac{1 - [G(t + xh(t))]^a \left\{ 1 + \lambda - \lambda [G(t + xh(t))]^b \right\}}{1 - [G(t)]^a \left\{ 1 + \lambda - \lambda [G(t)]^b \right\}} \\ &= \lim_{t \to w(F)} \frac{1 - [G(t + xh(t))]^a - \lambda [G(t + xh(t))]^a \left\{ 1 - [G(t + xh(t))]^b \right\}}{1 - [G(t)]^a - \lambda [G(t)]^a \left\{ 1 - [G(t + xh(t))]^b \right\}} \\ &= \lim_{t \to w(G)} \frac{1 - [G(t + xh(t))]^a - \lambda [G(t)]^a \left\{ 1 - [G(t + xh(t))]^b \right\}}{1 - [G(t)]^a - \lambda [G(t)]^a \left\{ 1 - [G(t + xh(t))]^b \right\}} \\ &= \lim_{t \to w(G)} \frac{1 - 1^a - \lambda \cdot 1^a \cdot \left\{ 1 - [G(t + xh(t))]^b \right\}}{1 - 1^a - \lambda \cdot 1^a \cdot \left\{ 1 - [G(t)]^b \right\}} \\ &= \lim_{t \to w(G)} \frac{1 - [G(t + xh(t))]^b}{1 - [G(t)]^b} \\ &= \lim_{t \to w(G)} \frac{1 - [G(t + xh(t))]^b}{1 - [G(t)]^b} \\ &= \lim_{t \to w(G)} \frac{1 - [1 - G(t + xh(t))]]^b}{1 - [1 - G(t + xh(t))]]^b} \\ &= \lim_{t \to w(G)} \frac{1 - [1 - G(t + xh(t))]}{1 - [1 - G(t)]^b} \\ &= \lim_{t \to w(G)} \frac{1 - [G(t + xh(t))]}{1 - [1 - G(t)]} \\ &= \lim_{t \to w(G)} \frac{1 - G(t + xh(t))}{1 - G(t)} \\ &= \exp(-x) \end{split}$$

for every x > 0, assuming w(F) = w(G). So, it follows that F also belongs to the max domain of attraction of the Gumbel extreme value distribution with

$$\lim_{n \to \infty} P\left(\frac{M_n - b_n}{a_n} \le x\right) = \exp\left[-\exp\left(-x\right)\right]$$

for some suitable norming constants $a_n > 0$ and b_n .

Second, suppose that G belongs to the max domain of attraction of the Fréchet extreme value distribution. Then, there must exist a $\beta > 0$ such that

$$\lim_{t \to \infty} \frac{1 - G(tx)}{1 - G(t)} = x^{-\beta}$$

for every x > 0. But

$$\begin{split} \lim_{t \to \infty} \frac{1 - F(tx)}{1 - F(t)} \\ &= \lim_{t \to \infty} \frac{1 - [G(tx)]^a \left\{ 1 + \lambda - \lambda [G(tx)]^b \right\}}{1 - [G(t)]^a \left\{ 1 + \lambda - \lambda [G(t)]^b \right\}} \\ &= \lim_{t \to \infty} \frac{1 - [G(tx)]^a - \lambda [G(tx)]^a \left\{ 1 - [G(tx)]^b \right\}}{1 - [G(t)]^a - \lambda [G(t)]^a \left\{ 1 - [G(tx)]^b \right\}} \\ &= \lim_{t \to \infty} \frac{1 - 1^a - \lambda \cdot 1^a \cdot \left\{ 1 - [G(tx)]^b \right\}}{1 - 1^a - \lambda \cdot 1^a \cdot \left\{ 1 - [G(tx)]^b \right\}} \\ &= \lim_{t \to \infty} \frac{1 - [G(tx)]^b}{1 - [G(t)]^b} \\ &= \lim_{t \to \infty} \frac{1 - [I - [1 - G(tx)]]^b}{1 - [I - G(tx)]^b} \\ &= \lim_{t \to \infty} \frac{1 - \{1 - [I - G(tx)]\}^b}{1 - \{1 - [I - G(tx)]\}^b} \\ &= \lim_{t \to \infty} \frac{1 - \{1 - [I - G(tx)]\}^b}{1 - \{1 - [I - G(t)]\}^b} \\ &= \lim_{t \to \infty} \frac{1 - [G(tx)]}{1 - [I - G(t)]} \\ &= \lim_{t \to \infty} \frac{1 - G(tx)}{1 - G(t)} \\ &= x^{-\beta}. \end{split}$$

for every x > 0. So, it follows that F also belongs to the max domain of attraction of the Fréchet extreme value distribution with

$$\lim_{n \to \infty} P\left(\frac{M_n - b_n}{a_n} \le x\right) = \exp\left(-x^{-\beta}\right)$$

for some suitable norming constants $a_n > 0$ and b_n .

Third, suppose that G belongs to the max domain of attraction of the Weibull extreme value distribution. Then, there must exist a $\beta > 0$ such that

$$\lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - t)} = x^{\beta}$$

for every x > 0. But

$$\begin{split} \lim_{t \to 0} \frac{1 - F(w(F) - tx)}{1 - F(w(F) - t)} \\ &= \lim_{t \to 0} \frac{1 - [G(w(F) - tx)]^a \left\{ 1 + \lambda - \lambda [G(w(F) - tx)]^b \right\}}{1 - [G(w(F) - t)]^a \left\{ 1 + \lambda - \lambda [G(w(F) - tx)]^b \right\}} \\ &= \lim_{t \to 0} \frac{1 - [G(w(F) - tx)]^a - \lambda [G(w(F) - tx)]^a \left\{ 1 - [G(w(F) - tx)]^b \right\}}{1 - [G(w(F) - t)]^a - \lambda [G(w(F) - t)]^a \left\{ 1 - [G(w(F) - t)]^b \right\}} \\ &= \lim_{t \to 0} \frac{1 - [G(w(G) - tx)]^a - \lambda [G(w(G) - tx)]^a \left\{ 1 - [G(w(G) - tx)]^b \right\}}{1 - [G(w(G) - t)]^a - \lambda [G(w(G) - tx)]^a \left\{ 1 - [G(w(G) - tx)]^b \right\}} \\ &= \lim_{t \to 0} \frac{1 - 1^a - \lambda \cdot 1^a \cdot \left\{ 1 - [G(w(G) - tx)]^b \right\}}{1 - 1^a - \lambda \cdot 1^a \cdot \left\{ 1 - [G(w(G) - tx)]^b \right\}} \\ &= \lim_{t \to 0} \frac{1 - [G(w(G) - tx)]^b}{1 - [G(w(G) - t)]^b} \\ &= \lim_{t \to 0} \frac{1 - [I - G(w(G) - tx)]^b}{1 - [I - G(w(G) - tx)]^b} \\ &= \lim_{t \to 0} \frac{1 - \left\{ 1 - [I - G(w(G) - tx)] \right\}^b}{1 - \left\{ 1 - [I - G(w(G) - tx)] \right\}} \\ &= \lim_{t \to 0} \frac{1 - \left\{ 1 - [I - G(w(G) - tx)] \right\}^b}{1 - \left\{ 1 - [G(w(G) - tx)] \right\}} \\ &= \lim_{t \to 0} \frac{1 - \left\{ 1 - [I - G(w(G) - tx)] \right\}^b}{1 - \left\{ 1 - G(w(G) - tx) \right\}} \\ &= \lim_{t \to 0} \frac{1 - \left\{ 1 - [I - G(w(G) - tx)] \right\}}{1 - \left\{ 1 - G(w(G) - tx) \right\}} \\ &= \lim_{t \to 0} \frac{1 - \left\{ 1 - [I - G(w(G) - tx)] \right\}}{1 - \left\{ 1 - G(w(G) - tx) \right\}} \\ &= \lim_{t \to 0} \frac{1 - \left\{ 1 - [G(w(G) - tx)] \right\}}{1 - \left\{ 1 - G(w(G) - tx) \right\}} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - t)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w(G) - tx)}{1 - G(w(G) - tx)} \\ \\ &= \lim_{t \to 0} \frac{1 - G(w($$

for every x > 0, assuming w(F) = w(G). So, it follows that F also belongs to the max domain of attraction of the Weibull extreme value distribution with

$$\lim_{n \to \infty} P\left(\frac{M_n - b_n}{a_n} \le x\right) = \exp\left(-(-x)^{\beta}\right)$$

for some suitable norming constants $a_n > 0$ and b_n .