

MATH4/68181: Extreme values and financial risk
Semester 1
Problem sheet for Week 10

1) Consider a bivariate distribution specified by the joint survival function

$$\bar{F}(x, y) = \exp \left[-\frac{\theta y^2}{x + y} + \theta y - x - y \right]$$

for $x > 0$ and $y > 0$.

- (a) show that the distribution is a bivariate extreme value distribution;
- (b) derive the joint cdf;
- (c) derive the conditional cdf if Y given $X = x$;
- (d) derive the conditional cdf if X given $Y = y$;
- (e) derive the joint pdf;
- (f) derive the conditional pdf of Y given $X = x$;
- (g) derive the conditional pdf of X given $Y = y$.

2) Consider a bivariate distribution specified by the joint survival function

$$\bar{F}(x, y) = \exp \left[\frac{\alpha xy}{x + y} - x - y \right]$$

for $x > 0$ and $y > 0$.

- (a) show that the distribution is a bivariate extreme value distribution;
- (b) derive the joint cdf;
- (c) derive the conditional cdf if Y given $X = x$;
- (d) derive the conditional cdf if X given $Y = y$;
- (e) derive the joint pdf;
- (f) derive the conditional pdf of Y given $X = x$;
- (g) derive the conditional pdf of X given $Y = y$.

3) Consider a bivariate distribution specified by the joint survival function

$$\bar{F}(x, y) = \exp \left[-(x^a + y^a)^{1/a} \right]$$

for $x > 0$ and $y > 0$.

- (a) show that the distribution is a bivariate extreme value distribution;
- (b) derive the joint cdf;
- (c) derive the conditional cdf if Y given $X = x$;
- (d) derive the conditional cdf if X given $Y = y$;
- (e) derive the joint pdf;
- (f) derive the conditional pdf of Y given $X = x$;
- (g) derive the conditional pdf of X given $Y = y$.

1)

$$\bar{G}(x, y) = e^{-\frac{\theta y^2}{x+y}} + \theta y - x - y$$

$$\begin{aligned} x &> 0 \\ y &> 0 \\ 0 &< \theta < 1 \end{aligned}$$

(a)

$$\begin{cases} \bar{G}(0, y) = e^{-y} \\ \bar{G}(x, 0) = e^{-x} \end{cases}$$

$$\Rightarrow \left. \begin{aligned} G_X(x) &= 1 - e^{-x} \\ G_Y(y) &= 1 - e^{-y} \end{aligned} \right\} \begin{array}{l} \text{both are} \\ \text{Exp}(1) \\ \text{marginals} \end{array}$$

$$\begin{aligned} \bar{G}(x, y) &= e^{-(x+y)} \left[\frac{\theta y^2}{(x+y)^2} - \frac{\theta y}{x+y} + 1 \right] \\ &= e^{-(x+y)} \left[\theta \left(\frac{y}{x+y} \right)^2 - \theta \left(\frac{y}{x+y} \right) + 1 \right] \\ &= e^{-(x+y)} A \left(\frac{y}{x+y} \right) \end{aligned}$$

where $A(w) = \theta w^2 - \theta w + 1$.

We need to show that $A(\cdot)$ satisfies

i) $A(0) = 1$

ii) $A(1) = 1$

iii) $\max(w, 1-w) \leq A(w) \leq 1 \quad \forall w$

$$\Leftrightarrow \begin{cases} (\alpha) A(w) \geq w \quad \forall w \\ (\beta) A(w) \geq 1-w \quad \forall w \\ (\gamma) A(w) \leq 1 \quad \forall w \end{cases}$$

iv) $A(\cdot)$ is convex.

$$(i) A(0) = \theta \cdot 0^2 - \theta \cdot 0 + 1 = 1 \checkmark$$

$$(ii) A(1) = \theta \cdot 1^2 - \theta \cdot 1 + 1 = 1 \checkmark$$

$$(x) A(w) \geq w \quad \forall w$$

$$\Leftrightarrow \theta w^2 - \theta w + 1 \geq w \quad \forall w$$

$$\Leftrightarrow \theta w(w-1) + 1-w \geq 0 \quad \forall w$$

$$\Leftrightarrow (1-\theta w)(1-w) \geq 0 \quad \forall w \checkmark$$

because $0 < \theta < 1$

$$(\beta) A(w) \geq 1-w \quad \forall w$$

$$\Leftrightarrow \theta w^2 - \theta w + 1 \geq 1-w \quad \forall w$$

$$\Leftrightarrow \theta w^2 + w(1-\theta) \geq 0 \quad \forall w \checkmark$$

because $0 < \theta < 1$

$$(\gamma) A(w) \leq 1 \quad \forall w$$

$$\Leftrightarrow \theta w^2 - \theta w + 1 \leq 1 \quad \forall w$$

$$\Leftrightarrow \theta w^2 - \theta w \leq 0 \quad \forall w$$

$$\Leftrightarrow \theta w(w-1) \leq 0 \quad \forall w \checkmark$$

$$(iv) A(w) = \theta w^2 - \theta w + 1$$

$$A'(w) = 2\theta w - \theta$$

$$A''(w) = 2\theta > 0$$

$\Rightarrow A(\cdot)$ is convex.

$\Rightarrow \bar{G}$ corresponds to a bivariate extreme value distribution.

(b) joint CDF

$$G(x, y) = 1 - \bar{G}(x, 0) - \bar{G}(0, y) + \bar{G}(x, y)$$

(c) conditional CDF of Y given $X=x$

$$\frac{\frac{\partial G(x, y)}{\partial y}}{e^{-x}}$$

(d) Conditional CDF of X given $Y=y$

$$\frac{\frac{\partial G(x, y)}{\partial x}}{e^{-y}}$$

(e) Joint PDF

$$\frac{\partial^2}{\partial x \partial y} G(x, y) = g(x, y)$$

(f) Conditional PDF of Y given $X=x$

$$\frac{g(x, y)}{e^{-x}}$$

(g) Conditional PDF of X given $Y=y$

$$\frac{g(x, y)}{e^{-y}}$$

2) a)

$$\bar{G}(x, y) = e^{-\frac{ax+y}{x+y} - x - y}$$

$$\begin{aligned} x &> 0 \\ y &> 0 \\ 0 &< a < 1 \end{aligned}$$

$$\bar{G}(0, y) = e^{-y}$$

$$\bar{G}(x, 0) = e^{-x}$$

$$\Rightarrow \left. \begin{aligned} G_X(x) &= 1 - e^{-x} \\ G_Y(y) &= 1 - e^{-y} \end{aligned} \right\} \begin{array}{l} \text{Exp}(1) \\ \text{marginals} \end{array}$$

$$\begin{aligned} \bar{G}(x, y) &= e^{-(x+y)} \left[-\frac{axy}{(x+y)^2} + 1 \right] \\ &= e^{-(x+y)} \left[-a \left(\frac{y}{x+y} \right) \left(1 - \frac{y}{x+y} \right) + 1 \right] \\ &= e^{-(x+y)} A \left(\frac{y}{x+y} \right) \end{aligned}$$

Where $A(w) = -aw(1-w) + 1$

(i) $A(0) = -a \cdot 0 \cdot (1-0) + 1 = 1 \checkmark$

(ii) $A(1) = -a \cdot 1 \cdot (1-1) + 1 = 1 \checkmark$

(A) $A(w) \geq w \quad \forall w$

$$\Leftrightarrow -aw(1-w) + 1 \geq w \quad \forall w$$

$$\Leftrightarrow (1-w)(1-aw) \geq 0 \quad \forall w \quad \checkmark$$

because $0 < a < 1$

(B) $A(w) \geq 1-w \quad \forall w$

$$\Leftrightarrow 1 - aw(1-w) \geq 1-w \quad \forall w$$

$$\Leftrightarrow w[1 - a(1-w)] \geq 0 \quad \forall w \quad \checkmark$$

because $0 < a < 1$

$$(i) \quad A(w) \leq 1 \quad \forall w$$

$$\Leftrightarrow 1 - \alpha w(1-w) \leq 1 \quad \forall w$$

$$\Leftrightarrow -\alpha w(1-w) \leq 0 \quad \forall w \quad \checkmark$$

$$(iv) \quad A(w) = 1 - \alpha w(1-w)$$

$$A'(w) = -\alpha + 2\alpha w$$

$$A''(w) = 2\alpha > 0$$

$\Rightarrow A(\cdot)$ is convex.

Hence, $\bar{G}(x, y)$ corresponds to a bivariate extreme value distribution. The remaining parts are similar to Q1.

$$3) \quad \bar{G}(x, y) = e^{-(x^a + y^a)^{\frac{1}{a}}}, \quad \begin{array}{l} a > 0 \\ x > 0 \\ y > 0 \end{array}$$

$$\bar{G}(0, y) = e^{-y}$$

$$\bar{G}(x, 0) = e^{-x}$$

$$\Rightarrow \left. \begin{array}{l} G_X(x) = 1 - e^{-x} \\ G_Y(y) = 1 - e^{-y} \end{array} \right\} \text{Exp}(1) \text{ marginals}$$

$$\begin{aligned} \bar{G}(x, y) &= e^{-(x+y) \left[\left(\frac{x}{x+y} \right)^a + \left(\frac{y}{x+y} \right)^a \right]^{\frac{1}{a}}} \\ &= e^{-(x+y) A\left(\frac{y}{x+y}\right)} \end{aligned}$$

$$\text{where } A(w) = \left[(1-w)^a + w^a \right]^{\frac{1}{a}}$$

$$(i) \quad A(0) = \left[(1-0)^a + 0^a \right]^{\frac{1}{a}} = 1 \quad \checkmark$$

$$(ii) \quad A(1) = \left[(1-1)^a + 1^a \right]^{\frac{1}{a}} = 1 \quad \checkmark$$

$$(iii) \quad A(w) \geq w \quad \forall w$$

$$\Leftrightarrow \left[w^a + (1-w)^a \right]^{\frac{1}{a}} \geq w \quad \forall w$$

$$\Leftrightarrow \cancel{w^a} + (1-w)^a \geq \cancel{w^a} \quad \forall w$$

$$\Leftrightarrow (1-w)^a \geq 0 \quad \forall w \quad \checkmark$$

$$(B) \quad A(w) \geq 1-w \quad \forall w$$

$$\Leftrightarrow \left[w^a + (1-w)^a \right]^{\frac{1}{a}} \geq 1-w \quad \forall w$$

$$\Leftrightarrow w^a + \cancel{(1-w)^a} \geq \cancel{(1-w)^a} \quad \forall w$$

$$\Leftrightarrow w^a \geq 0 \quad \forall w \quad \checkmark$$

$$\begin{aligned}
 (\gamma) \quad & \omega \leq 1 \\
 \Rightarrow & \omega^a \leq \omega \quad \text{assuming } \underline{a \geq 1} \\
 \Rightarrow & (1-\omega)^a \leq 1-\omega \quad \parallel \quad \parallel \\
 \Rightarrow & \omega^a + (1-\omega)^a \leq 1 \quad \forall \omega \\
 \Rightarrow & [\omega^a + (1-\omega)^a]^{\frac{1}{a}} \leq 1 \quad \forall \omega \\
 \Rightarrow & A(\omega) \leq 1 \quad \forall \omega
 \end{aligned}$$

$$\begin{aligned}
 (\text{iv}) \quad & A(\omega) = [\omega^a + (1-\omega)^a]^{\frac{1}{a}} \\
 & A'(\omega) = [\omega^a + (1-\omega)^a]^{\frac{1}{a}-1} (\omega^{a-1} - (1-\omega)^{a-1}) \\
 & A''(\omega) = (a-1) [\omega^a + (1-\omega)^a]^{\frac{1}{a}-2} > 0 \\
 & \cdot \left\{ \underbrace{[\omega^{a-1} - (1-\omega)^{a-1}]^2}_{> 0} + \underbrace{[\omega^a + (1-\omega)^a]}_{> 0} \underbrace{\left[\frac{a-2}{\omega} - \frac{a-2}{1-\omega} \right]}_{> 0} \right\} \\
 & > 0 \quad \forall \omega
 \end{aligned}$$

$\Rightarrow A(\cdot)$ is convex.

Hence, \bar{G} corresponds to a bivariate extreme value distribution.

The remaining parts are similar to $\mathcal{A}1$.